

G. The perfect matching function

References:

- Alexander A. Razborov, A lower bound on the monotone network complexity of the logical permanent, *Math. Notes Acad. Sci. USSR* 37 (1985), 485 - 493.
- Stasys Jukna, *Boolean Function Complexity: Advances and Frontiers*, Springer 2012, 274 - 281.

The perfect matching function PM_m is a monotone Boolean function of $n := m^2$ variables which correspond to the edge set of a bipartite graph $G = (A, B, E)$ where $|A| = |B| = m$. $PM_m(x) = 1$ iff the corresponding graph G contains a perfect matching; i.e., a set of m vertex-disjoint edges. Since a perfect matching in a given bipartite graph can be computed in polynomial time, the non-monotone complexity of PM_m is also polynomial.

In 1985, Alexander Razborov has constructed approximators to prove an $m^{2c \log m}$ lower bound for the monotone network complexity of PM_m . As mentioned in their paper, Berg and Ulfberg have investigated the construction of DNF/CNF - approximators to prove a super-polynomial lower bound for the monotone net =

work complexity of PM_m .

Question:

Why Razborov-approximators suffice but DNF/CNF-approximators do not suffice to prove a super-polynomial lower bound for the monotone complexity of PM_m ?

To get the possibility to answer this question, we shall investigate Razborov's proof.

Let $V = \{1, 2, \dots, m\}$. S_m denotes the set of all $m!$ permutations of $1, 2, \dots, m$. Then we can write

$$PM_m(x) = \bigvee_{\sigma \in S_m} \bigwedge_{i=1}^m x_{i\sigma(i)}$$

Goal:

Construction of a legitimate lattice $\mathcal{L} = (S, \tau, \rho)$ with large distance $\rho(CPM_m, S)$.

Let

$$Per_e := \{ E' \subseteq A \times B \mid E' \text{ is a matching of size } \leq e \}$$

Now we define \mathcal{L} exactly as above for the clique function where Per_e plays the role of VCE .

Lemma 6.1

$S = (S, r, l)$ constructed above is a legitimate lattice.

Proof:

Exercise

Let $G_+ = (A, B, E_+)$ be a random graph taking its values in the set of all $m!$ perfect matchings with equal probability $\frac{1}{m!}$.

Definition of $G_+ \Rightarrow$

$$\text{Prob}[PM_m(G_+) = 1] = 1.$$

Let $h: A \cup B \rightarrow \{0, 1\}$ be a random colouring assigning each node in $A \cup B$ a value 0 or 1 independently with probability $1/2$.

This colouring defines a random graph $G_- = (V, E_-)$ where

$$V = A \cup B \text{ and } E_- = \{(v, w) \mid h(v) = h(w)\}.$$

Note that in most cases, G_- is not a bipartite graph.

Lemma 6.2

$$\text{Prob}[PM_m(G_-) = 0] \geq 1 - \frac{2}{m}.$$

Proof:

The graph $G_- = (V, E_-)$ contains a perfect matching iff

$$\sum_{v \in A} h(v) = \sum_{w \in B} h(w). \quad (\text{show this!})$$

\Rightarrow

$$\text{Prob} [PM_m(G_-) = 1] = \text{Prob} \left[\sum_{v \in A} h(v) = \sum_{w \in B} h(w) \right]$$

$$= \sum_{j=0}^m \text{Prob} \left[\sum_{v \in A} h(v) = j \right] \cdot \text{Prob} \left[\sum_{w \in B} h(w) = j \right]$$

$$\leq \max_{0 \leq j \leq m} \text{Prob} \left[\sum_{v \in A} h(v) = j \right]$$

$$\leq \binom{m}{\frac{m}{2}} \cdot 2^{-m} \leq \frac{2}{\sqrt{m}}$$

The last inequality is obtained using the Stirling formula.

Hence,

$$\text{Prob} [PM_m(G_-) = 0]$$

$$= 1 - \text{Prob} [PM_m(G_-) = 1]$$

$$\geq 1 - \frac{2}{\sqrt{m}}.$$

Given any two members M and N of the lattice S , our goal is to show that the

probabilities $\text{Prob}[G_+ \in \delta_n(N, N)]$ and $\text{Prob}[G_- \in \delta_n(M, N)]$ are small. We shall consider $\text{Prob}[G_+ \in \delta_n(N, M)]$ first.

Remember that G_+ is a random perfect matching.

Lemma 6.3

For any $M, N \in \mathcal{S}$ there holds

$$\text{Prob}[G_+ \in \delta_n(M, N)] \leq (r-1)^{2r} \frac{(m-2-1)!}{m!}$$

Proof:

Let $M = \Gamma A \Gamma$ and $N = \Gamma B \Gamma$ where $A, B \subseteq \text{Per}_r$ closed. Then

$$\begin{aligned} \delta_n(M, N) &= (\Gamma A \Gamma \cap \Gamma B \Gamma) \setminus \Gamma A \Gamma \cap \Gamma B \Gamma \\ &= (\Gamma A \Gamma \cap \Gamma B \Gamma) \setminus \Gamma A \cap B \Gamma. \end{aligned}$$

Let $A_m \subset A$ and $B_m \subset B$ denote the sets of minimal elements in A and B , respectively. Then

$$\begin{aligned} \delta_n(M, N) &= (\Gamma A_m \Gamma \cap \Gamma B_m \Gamma) \setminus \Gamma A \cap B \Gamma \\ &= \left(\bigcup_{E_1 \in A_m} \Gamma E_1 \Gamma \cap \bigcup_{E_2 \in B_m} \Gamma E_2 \Gamma \right) \setminus \Gamma A \cap B \Gamma \end{aligned}$$

$$= \bigcup_{E_1 \in A_m} \bigcup_{E_2 \in B_m} (\Gamma E_1 \Gamma \cap \Gamma E_2 \Gamma) \setminus \Gamma A \cap B \Gamma$$

Since $\Gamma E_1 \Gamma \cap \Gamma E_2 \Gamma = \Gamma E_1 \cup E_2 \Gamma$, we obtain

$$= \bigcup_{E_1 \in A_m} \bigcup_{E_2 \in B_m} (\{E_1 \cup E_2\} \setminus \{\Gamma \cap B\})$$

Consider any $E_1 \in A_m, E_2 \in B_m$. Then three cases are possible

Case 1: $E_1 \cup E_2$ is not a matching.

$$\text{Then } \text{Prob}[G_+ \in \{E_1 \cup E_2\}] = 0$$

Case 2: $E_1 \cup E_2$ is a matching and $|E_1 \cup E_2| \leq \ell$.

Then $E_1 \cup E_2 \in \text{Per}_\ell \Rightarrow$

- A closed and $E_1 \in A \Rightarrow E_1 \cup E_2 \in A$
- B closed and $E_2 \in B \Rightarrow E_1 \cup E_2 \in B$.

Hence, $E_1 \cup E_2 \in A \cap B$.

$$\Rightarrow \{E_1 \cup E_2\} \setminus \{\Gamma \cap B\} = \emptyset.$$

Case 3: $E_1 \cup E_2$ is a matching but $|E_1 \cup E_2| > \ell$.

$$\begin{aligned} \text{Prob}[G_+ \in \{E_1 \cup E_2\}] &= \text{Prob}[E_1 \cup E_2 \subseteq E_+] \\ &= \frac{(m - |E_1 \cup E_2|)!}{m!} \\ &\leq \frac{(m - \ell - 1)!}{m!} \end{aligned}$$

Altogether, applying Lemma 4.3, we obtain

$$\begin{aligned} \text{Prob}[G_+ \in \mathcal{S}_\pi(M, N)] &\leq |A_m| \cdot |B_m| \cdot \frac{(m - \ell - 1)!}{m!} \\ &\leq (r-1)^{2\ell} \cdot \frac{(m - \ell - 1)!}{m!} \end{aligned}$$

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Since the events $e_1 \in E_-$ and $e_2 \in E_-$ are not necessarily independent, the proof of an upper bound for the probability $\text{Prob}(G_- \in \mathcal{S}_U(n, N))$ is more difficult. The following lemma shows that these events are independent if the edges come from a fixed forest.

Lemma 6.4

Let $E' := \{(v_1, w_1), (v_2, w_2), \dots, (v_p, w_p)\} \subset A \times B$ be a forest F . Then the events $(v_i, w_i) \in E_-$ are independent, and each happens with probability $\frac{1}{2}$.

Proof.

Let $h: A \cup B \rightarrow \{0, 1\}$ be a random colouring assigning each node in $A \cup B$ a value 0 or 1 independently with probability $\frac{1}{2}$.

Assume that h is performed in such an order that for each node u unequal the root of a tree in F its father has been coloured before.

For any edge $(v_i, w_i) \in E'$ consider the moment when the second node of (v_i, w_i) is coloured. Then

$$(v_i, w_i) \in E_- \iff h(v_i) = h(w_i)$$

This occurs with probability $\frac{1}{2}$ and do not depend on the fact if some other edges are in E_- or not.

The following lemma shows that any subset $M \subseteq \text{Per}_e$ of r pairwise disjoint matchings contains a large subset $M_0 \subseteq M$ such that the union of the matchings in M_0 is a forest.

Lemma 6.5

Let $M \subseteq \text{Per}_e$ be a set of $|M| = r$ pairwise disjoint matchings. Then there exists a subset $M_0 \subseteq M$ of $|M_0| \geq \frac{\sqrt{r}}{2}$ matchings such that $\bigcup_{A \in M_0} A$ is a forest.

Proof:

Let $M_0 \subseteq M$ such that $\bigcup_{A \in M_0} A$ is a forest and $|M_0|$ is maximal.

It suffices to show that $|M_0| \geq \frac{\sqrt{r}}{2}$.

Assume that $|M_0| < \frac{\sqrt{r}}{2}$. Let $E_0 := \bigcup_{A \in M_0} A$.

$\Rightarrow |E_0| < \sqrt{r}$.

Let $A_0 \subseteq A$ and $B_0 \subseteq B$ be the sets of nodes which are end node of at least one edge in E_0 .

$\Rightarrow |A_0| < \sqrt{r}$ and $|B_0| < \sqrt{r}$.

Since M contains $|M| = r > |A_0 \times A_0|$ pairwise disjoint matchings at least one of

these matchings, say A_1 , contains no edge in $A_0 \times B_0$. (186)

Since A_1 is a matching and E_0 is a forest lying in $A_0 \times B_0$, the graph $E_0 \cup A_1$ is a forest as well.

But $A_1 \cap E_0 = \emptyset$ implies $A_1 \notin M_0$, a contradiction to the maximality of M_0 . ■

Note that

$$|\text{Per}_e| \leq \sum_{i=0}^{\ell} \binom{m}{i}^2 \cdot i! \leq m^e \sum_{i=0}^{\ell} \binom{m}{i} \leq m^{2\ell}$$

Now we are prepared to prove an upper bound for $\text{Prob}[G \in \delta_{\square}(M, N)]$

Lemma 6.6

For any $M, N \in \mathcal{S}$, there holds

$$\text{Prob}[G \in \delta_{\square}(M, N)] \leq (1 - 2^{-r})^{\frac{r}{k}} \cdot 2r^{\ell}$$

Proof:

Let $M = \Gamma A$, $N = \Gamma B$ and $C = A \cup B$.

\Rightarrow

$$\delta_{\square}(M, N) = \Gamma C^* \setminus C.$$

Let $C = C_0, C_1, C_2, \dots, C_p = C^*$ be the results of the improvement steps in the construction of C^* from C .

Let $W_i \in \{W \in C_{i-1} \mid C_{i-1} \vdash W\}$ be the chosen set for the construction of C_i from C_{i-1} .

Since $p \leq 2r^e$ it suffices to prove that

$$(*) \text{ Prob}[G_- \in \Gamma C_i \setminus \Gamma C_{i-1}] \leq (1 - 2^{-e})^{\frac{1-p}{e}}$$

Let $A_1, A_2, \dots, A_r \in C_{i-1} : A_1, A_2, \dots, A_r \perp W$.

Definition of $\perp \Rightarrow$

$A_1 \perp W, A_2 \perp W, \dots, A_r \perp W$ are pairwise disjoint matchings in Per_e .

If at least one of these matchings is empty then a subset of W would be contained in C_{i-1} and hence, $\Gamma W \subseteq \Gamma C_{i-1}$. By construction, $\Gamma C_i = \Gamma C_{i-1}$ such that $(*)$ trivially holds.

Otherwise, applying Lemma 6.5, we choose a subset $M_0 \subseteq \{A_1 \perp W, A_2 \perp W, \dots, A_r \perp W\}$ such that $\bigcup_{A \in M_0} A$ is a forest and $|M_0| \geq \frac{1-p}{e}$.

\Rightarrow

$$\text{Prob}[G_- \in \Gamma C_i \setminus \Gamma C_{i-1}]$$

$$\leq \text{Prob}[W \subseteq E_- \text{ and } A_i \not\subseteq E_- \forall i=1, \dots, r]$$

$$\leq \text{Prob}[A_i \perp W \not\subseteq E_- \forall i=1, \dots, r]$$

$$\leq \text{Prob}[A_i \perp W \not\subseteq E_- \forall A_i \perp W \in M_0]$$

Lemma 6.4 \Rightarrow

All events $A_i \perp W \subseteq E_-$ for $A_i \perp W \in M_0$ are independent and

$$\text{Prob}[A_i \perp W \subseteq E_-] = 2^{-|A_i \perp W|} \geq 2^{-e}$$

Therefore,

$$\text{Pr} [A_i | W \in E_- \wedge A_i | W \in M_0]$$

$$= \prod_{A_i | W \in M_0} \text{Pr} [A_i | W \notin E_-]$$

$$\leq (1 - 2^{-r})^{\frac{T}{e}}$$

Theorem 6.1

$$C_{\sqrt{2m}}(PM_m) \gg m^{\sqrt{2}(\log m)}$$

Proof:

Let $t = g(PM_m, S)$ where

$$L := \lfloor \frac{1}{2} \log m \rfloor \text{ and } r := \lfloor m^{1/4} \log^2 m \rfloor.$$

Considers $M_1, N_1, M_2, N_2, \dots, M_t, N_t \in S$
such that

$$\sigma(PM_m) \subseteq M \cup \bigcup_{i=1}^t \delta_{\Pi}(M_i, N_i)$$

and

$$M \subseteq \sigma(PM_m) \cup \bigcup_{i=1}^t \delta_{\cup}(M_i, N_i).$$

We distinguish two cases.

Case 1: $M = \emptyset$.

$$\text{Then } \sigma(\emptyset) \subseteq \bigcup_{i=1}^t \delta_{\Pi}(M_i, N_i).$$

Since $G_+ \in \sigma(PM_m)$ with probability 1, the sum of probabilities $\text{Pr}(G_+ \in \delta_{\Pi}(M_i, N_i))$

at least 1 as well. Hence, by Lemma 6.3

$$t \cdot (r-1)^{2e} \frac{(m-e-1)!}{m!} \geq 1$$

$$\Leftrightarrow t \geq \frac{m!}{(m-e-1)! (r-1)^{2e}}$$

$$\geq \left(\frac{m}{2}\right)^e r^{-2e}$$

$$= \left(\frac{m}{2r^2}\right)^e$$

$$\geq \left(\frac{m}{2(m^{1/2} \log^8 m)}\right)^{\frac{\log m}{8}}$$

$$= m^{\Omega(\log m)}$$

Case 2: $M \neq \emptyset$.

Then $\exists A \in \text{Per}_e$ such that $\Gamma A \subseteq M$

\Rightarrow

$$\Gamma A \subseteq \text{OT}(\text{PM}_m) \cup \bigcup_{i=1}^t \delta_{\square}(M_i, N_i)$$

Lemma 6.4 \Rightarrow

$$\text{Prob}(G_e \in \Gamma A) = 2^{-|A|} \geq 2^{-e}$$

Lemma 6.2 \Rightarrow

$$\text{Prob}(G_e \in \text{OT}(\text{PM}_m)) \leq 2 \cdot m^{-1/2}$$

Lemma 6.6 \Rightarrow

$$\begin{aligned} \text{Prob}(G_e \in \delta_{\square}(M_i, N_i)) &\leq (1-2^{-e})^{\frac{1}{2}} \cdot 2r^e \\ &\leq (1-2^{-e})^{\frac{1}{2}} \cdot m^{2e} \end{aligned}$$

Hence we obtain

$$t \geq (2^{-\epsilon} - 2m^{-1/2}) (1 - 2^{-\epsilon})^{-\frac{\sqrt{t}}{2}} m^{-2\epsilon}$$

Since $1 - x \geq e^{-x - \frac{x^2}{2}}$ for $0 < x < 1$ we obtain

$$t \geq (2^{-\epsilon} - 2m^{-1/2}) e^{2^{-\epsilon} \frac{\sqrt{t}}{2} + \frac{2^{-2\epsilon}}{2} \cdot \frac{\sqrt{t}}{2}} \cdot m^{-2\epsilon}$$

$$\geq (2^{-\epsilon} - 2m^{-1/2}) e^{2^{-\epsilon} \cdot \frac{\sqrt{t}}{2}} \cdot m^{-2\epsilon}$$

$$\geq \frac{1}{8} m^{-1/8} \cdot e^{2^{-\epsilon} \cdot \frac{\sqrt{t}}{2}}$$

$$\geq \frac{1}{8} m^{-1/8 - 1/4 \log m} \cdot e^{\frac{m^{-2\epsilon} \cdot m^{-1/8} \cdot m^{1/8} \cdot \log^4 m}{\log m}}$$

$$= m^{-\Omega(\log^2 m)}$$

This proves the theorem.

Let us return to our former question:

Why Razborov - approximators suffice but DNF/CNF - approximators do not suffice to prove a super-polynomial lower bound for the monotone complexity of PM_m ?

Knowing Razborov's solution, it is still difficult to answer the above question.

First, one could try to use Razborov's test sets and his arguments to prove the needed lemmas with respect to DNF/CNF - approximators.

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